Scoring anomalies among multivariate extreme observations

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Plan



- 2 Multivariate EVT & Extreme Dependence
- 3 Sparse dependence structure

4 Estimation

5 Experiments

Anomaly Detection (AD)

What is Anomaly Detection ?

"Finding patterns in the data that do not conform to expected behavior"



Machine Learning context

Different kind of Anomaly Detection

Supervised AD

- Labels available for both normal data and anomalies
- Similar to rare class mining

Semi-supervised AD

- Only normal data available to train
- The algorithm learns on normal data only

Unsupervised AD

- no labels, training set = normal + abnormal data
- Assumption: anomalies are very rare

Anomaly Detection Schemes



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• Step 1: Learn a profile of the "normal" behavior Profile can be patterns, summary statistics,...

Anomaly Detection Schemes



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- Step 2: Use the "normal" profile to build a decision function.

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Anomaly Detection Schemes



- Step 1: Learn a profile of the "normal" behavior Profile can be patterns, summary statistics,...
- Step 2: Use the "normal" profile to build a decision function.
- **Step 3: Detect anomalies** among new observations. Anomalies are observations whose characteristics differ significantly from the normal profile

General idea of our work

- Extreme observations play a special role when dealing with outlying data.
- But no algorithm has specific treatment for such multivariate extreme observations.
- Our goal: Provide a method which can improve performance of standard AD algorithms by combining them with a **multivariate extreme analysis** of the **dependence structure**.

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Towards high dimension

Main idea :

Sparsity Assumption on the dependence structure:

Small number ($<<< 2^d$) of "groups of coordinates which can be large together"

Our Goal:

Learn the sparsity pattern of the dependence structure

Applications:

- Detecting Anomalies with Multivariate Extremes (DAMEX): anomalies = points violating the sparsity pattern.
- **Dimension reduction** If furthermore only a moderate number of *X_j*'s may be large together. Preliminary step before inference of the joint distribution of extremes.





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Framework

Context

- Random vector $\mathbf{X} = (X_1, \dots, X_d)$
- Margins: $X_j \sim F_j$ (F_j continuous)

Preliminary step: Standardization of each marginal

Standard Pareto:
$$V_j = \frac{1}{1 - F_j(X_j)}$$
 $\left(\mathbb{P}(V_j \ge x) = \frac{1}{x}, x \ge 1 \right)$

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Problematic

Joint extremes: V's distribution above large thresholds?

 $\mathbb{P}(\mathbf{V} \in \mathbf{A})$? (**A** 'far from the origin').



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Fundamental hypothesis and consequences

• Standard assumption: let A extreme region,

 $\mathbb{P}[\mathbf{V} \in t \ A] \simeq t^{-1} \mathbb{P}[\mathbf{V} \in A]$ (radial homogeneity)

• Formally,

regular variation (after standardization):

0 ∉ **A**

$$t\mathbb{P}[\mathbf{V} \in t \ A] \xrightarrow[t \to \infty]{} \mu(A), \qquad \mu : \text{ exponent measure}$$

Necessarily: $\mu(tA) = t^{-1}\mu(A)$

• \Rightarrow angular measure on sphere S_{d-1} : $\Phi(B) = \mu\{tB, t \ge 1\}$

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General model in multivariate EVT

Model for excesses

For an extreme region A:

$$\mathbb{P}[\mathbf{V} \in \mathbf{A}] \simeq \mu(\mathbf{A})$$

 \Leftrightarrow For a large r > 0 and a region *B* on the unit sphere:

$$\mathbb{P}\left[\|\mathbf{V}\| > r, \ \frac{\mathbf{V}}{\|\mathbf{V}\|} \in B\right] \simeq \frac{1}{r} \Phi(B)$$

 $\Rightarrow \Phi$ (or $\mu)$ rules the joint distribution of extremes (if margins are known).

Angular distribution

• Φ rules the joint distribution of extremes



• Asymptotic dependence: (V_1, V_2) may be large together.

vs

• Asymptotic independence: only V_1 or V_2 may be large.

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Dependence structure



Full pattern : anything may happen



Sparse pattern $(V_1 \text{ not large if } V_2 \text{ or } V_3 \text{ large})$

Sub-cones:

$$\mathcal{C}_{\alpha} = \left\{ \| \mathbf{v} \| \geq 1, \ \mathbf{v}_i > \mathbf{0} \ (i \in \alpha), \ \mathbf{v}_j = \mathbf{0} \ (j \notin \alpha) \right\}$$

• Corresponding sub-spheres: $\{\Omega_{\alpha}, \alpha \subset \{1, \dots, d\}\}$ $(\Omega_{\alpha} = C_{\alpha} \cap S_{d-1})$

Representation of extreme data

$$\mathcal{M} = \left\{ \begin{array}{ll} \Phi(\Omega_{\alpha}) : & \emptyset \neq \alpha \subset \{1, \ \dots, \ d\} \end{array} \right\} \\ &= \left\{ \begin{array}{ll} \mu(\mathcal{C}_{\alpha}) : & \emptyset \neq \alpha \subset \{1, \ \dots, \ d\} \end{array} \right\}$$



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Assumption: $\frac{d\Phi_{|\Omega_{\alpha}}}{dv_{\alpha}} = O(1).$

Representation $\mathcal M$ is natural and linear (after non-linear transform of the data $X \to V).$





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Problem: \mathcal{M} is an **asymptotic** representation

$$\mathcal{M} \ = \ \left\{ \ \Phi(\Omega_\alpha), \ \alpha \ \right\} \ = \ \left\{ \ \mu(\mathcal{C}_\alpha), \ \alpha \ \right\}$$

is the restriction of an asymptotic measure

$$\mu(A) = \lim_{t \to \infty} t \mathbb{P}[\mathbf{V} \in t \ A]$$

to a representative class of set $\{C_{\alpha}, \alpha\}$, but only the central sub-cone has positive Lebesgue measure!



Solution

Fix $\varepsilon > 0.$ Affect data $\varepsilon\text{-close}$ to an edge, to that edge.



$$\begin{split} \Omega_{\alpha} &\to \Omega_{\alpha}^{\epsilon} = \{ \boldsymbol{v} \in \mathbf{S}_{d-1} : \boldsymbol{v}_{j} > \epsilon \; (i \in \alpha), \; \boldsymbol{v}_{j} \leq \epsilon \; (j \notin \alpha) \}. \\ \mathcal{C}_{\alpha} &\to \mathcal{C}_{\alpha}^{\epsilon} = \{ t \; \Omega_{\alpha}^{\epsilon}, t \geq 1 \} \end{split}$$

New partition of S_{d-1} , compatible with non asymptotic data.

$$\hat{V}_{i}^{j} = \frac{1}{1 - \hat{F}_{j}(X_{i}^{j})}$$
 with $\hat{F}_{j}(X_{i}^{j}) = \frac{\operatorname{rank}(X_{i}^{j}) - 1}{n}$

 \Rightarrow get an natural estimate of $\Phi(\Omega_{\alpha})$

$$\begin{split} \widehat{\Phi}(\Omega_{\alpha}) &:= \frac{n}{k} \mathbb{P}_{n}(\widehat{V} \in \frac{n}{k} \mathcal{C}_{\alpha}^{\epsilon}) \\ &(\frac{n}{k} \text{ large, } \epsilon \text{ small}) \end{split}$$



 \Rightarrow we obtain

$$\widehat{\mathcal{M}} := \big\{ \ \widehat{\Phi}(\Omega_{\alpha}), \ \alpha \ \big\}$$

Theorem:

$$\|\widehat{\mathcal{M}} - \mathcal{M}\|_{\infty} \leq Cd\left(\sqrt{\frac{1}{\epsilon k}\log\frac{d}{\delta}} + Md\epsilon\right) + bias(\epsilon, k, n),$$

Algorithm

DAMEX in O(dn log n)

Input: parameters $\epsilon > 0$, k = k(n), $\Phi_{\min} \ge 0$.

1 Standardize via marginal rank-transformation: $\hat{V}_i := (1/(1 - \hat{F}_j(X_i^j)))_{j=1,...,d}$.

- 2 Assign to each \hat{V}_i the cone C^{ε}_{α} it belongs to.
- **3** Compute $\Phi_n^{\alpha,\epsilon} := \widehat{\Phi}(\Omega_{\alpha}) = \frac{n}{k} \mathbb{P}_n(\widehat{V} \in \frac{n}{k} \mathcal{C}_{\alpha}^{\epsilon})$ the estimate of the α -mass of Φ .

Set to 0 the Φ^{α,ε}_n below some small threshold Φ_{min} ≥ 0 to eliminate cones with negligible mass

Output: (sparse) representation of the dependence structure

$$\widehat{\mathcal{M}} := (\Phi_n^{\alpha,\epsilon})_{\alpha \subset \{1,\dots,d\}, \Phi_n^{\alpha,\epsilon} > \Phi_{\min}}$$

Sparse low-dimensional representation?

(Shell Research, and thanks J. Wadsworth)

data=50 wave direction from buoys in North sea.



nb of faces with positive mass after 2nd thresholding

Application to Anomaly Detection

After standardization of marginals: $\mathbb{P}[R > r, W \in B] \simeq \frac{1}{r} \Phi(B)$

$$\rightarrow \text{ scoring function} = \Phi_n^{\epsilon} \times 1/r : \\ \boldsymbol{s}_n(\mathbf{x}) := (1/\|\hat{\boldsymbol{\tau}}(\mathbf{x})\|_{\infty}) \sum_{\alpha} \Phi_n^{\alpha,\epsilon} \mathbb{1}_{\hat{\boldsymbol{\tau}}(\mathbf{x}) \in \mathcal{C}_{\alpha}^{\epsilon}}.$$

where $T: \mathbf{X} \mapsto \mathbf{V}$ $(V_j = \frac{1}{1 - F_j(X_j)})$



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	number of samples	number of features
shuttle	85849	9
forestcover	286048	54
SA	976158	41
SF	699691	4
http	619052	3
smtp	95373	3

Table: Datasets characteristics

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Figure: ROC and PR curve on forestcover dataset



Figure: ROC and PR curve on http dataset



Figure: ROC and PR curve on smtp dataset



Figure: ROC and PR curve on SF dataset



Figure: ROC and PR curve on shuttle dataset



Figure: ROC and PR curve on SA dataset

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