

# Anomaly Detection in Scikit-Learn and new tools from Multivariate Extreme Value Theory

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Supervision:

Detecting Anomalies with Multivariate Extremes:  
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Contributions to Scikit-Learn: Alexandre Gramfort

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- 1 Anomaly Detection and Scikit-Learn
- 2 Multivariate EVT & Representation of Extremes
- 3 Estimation
- 4 Experiments

# Anomaly Detection (AD)

## What is Anomaly Detection ?

"Finding patterns in the data that do not conform to expected behavior"



Huge number of applications: Network intrusions, credit card fraud detection, insurance, finance, military surveillance,...

# Machine Learning context

## Different kind of Anomaly Detection

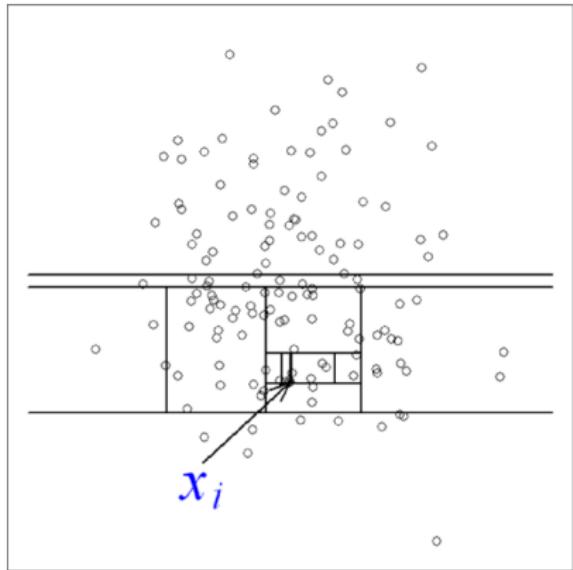
- **Supervised AD**
  - Labels available for both normal data and anomalies
  - Similar to rare class mining
- **Semi-supervised AD (Novelty Detection)**
  - Only normal data available to train
  - The algorithm learns on normal data only
- **Unsupervised AD (Outlier Detection)**
  - no labels, training set = normal + abnormal data
  - Assumption: anomalies are very rare

## Important literature in Anomaly Detection:

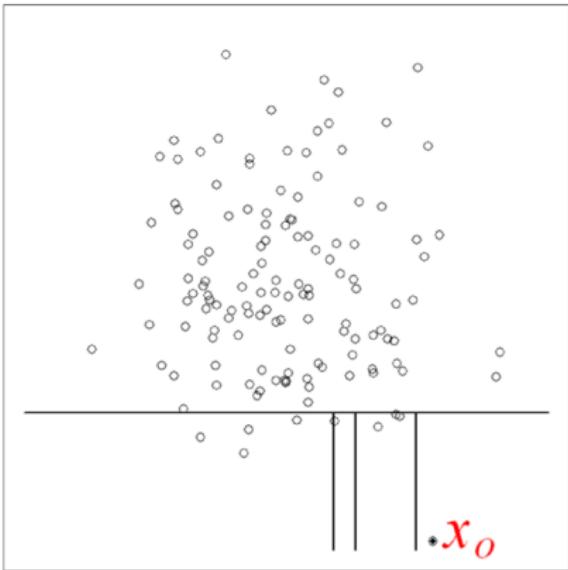
- **statistical AD techniques**  
fit a statistical model for normal behavior  
ex: [EllipticEnvelope](#)
- **density-based**  
- ex: Local Outlier Factor (LOF) and variants (COF ODIN LOCI)
- **Support estimation** - [OneClassSVM](#) - MV-set estimate
- **high-dimensional techniques:** - Spectral Techniques - Random Forest - [Isolation Forest](#)

# Isolation Forest:

Liu Tink Zhou icdm2008

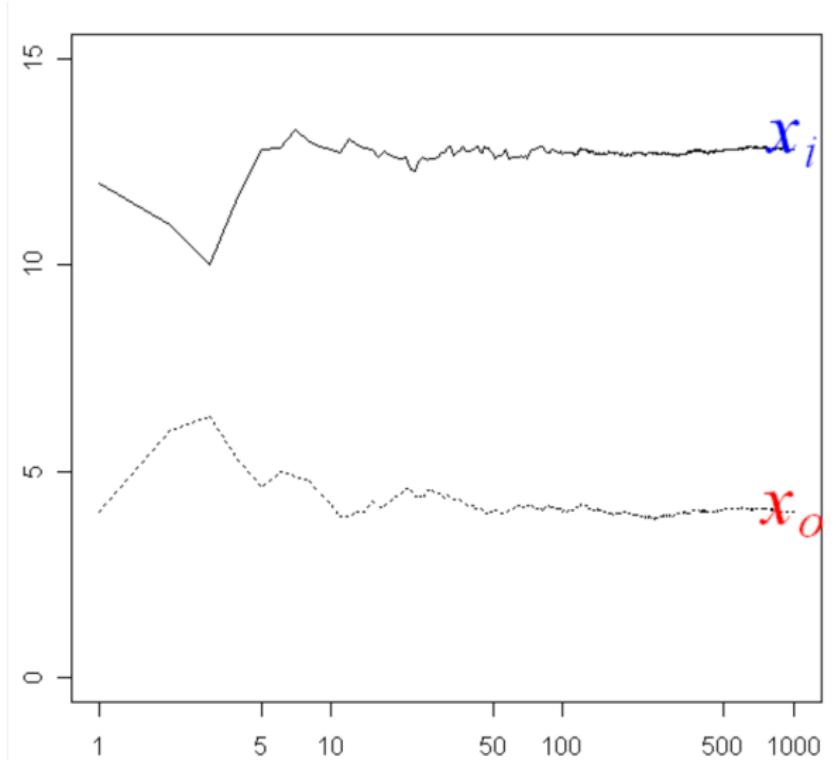


(a) Isolating  $x_i$



(b) Isolating  $x_o$

average path length



nb. of tree (log scale)

## IsolationForest.fit(X)

### IsolationForest

**Inputs:** X, n\_estimators, max\_samples

**Output:** Forest with:

- # trees = n\_estimators
- sub-sampling size = max\_samples
- maximal depth  $\max\_depth = \text{int}(\log_2 \max\_samples)$

Complexity:  $O(n_{\text{estimators}} \max_{\text{samples}} \log(\max_{\text{samples}}))$

default: n\_estimators=100, max\_samples=256

## IsolationForest.predict(X)

### Finding the depth in each tree

```
depth(Tree, X):  
    # – Finds the depth level of the leaf node  
    # for each sample x in X.  
    # – Add average_path_length(n_samples_in_leaf)  
    # if x not isolated
```

$$\text{score}(x, n) = 2^{-\frac{E(\text{depth}(x))}{c(n)}}$$

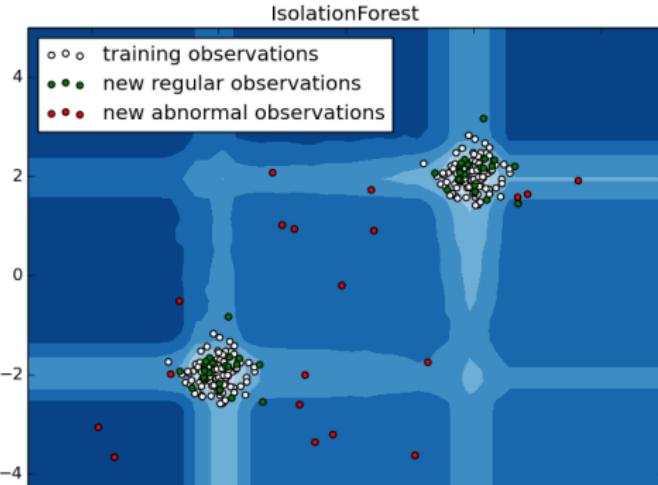
Complexity:  $O(n_{\text{samples}} n_{\text{estimators}} \log(\max_{\text{samples}}))$

# Examples

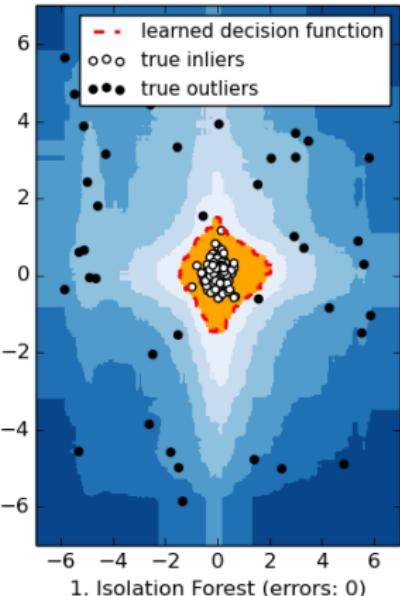
- code example:

```
>> from sklearn.ensemble import IsolationForest  
>> IF = IsolationForest()  
>> IF.fit(X_train) # build the trees  
>> IF.predict(X_test) # find the average depth
```

- plotting decision function:

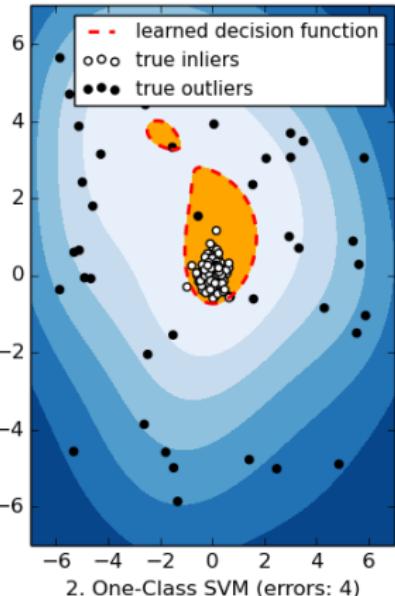


Outlier detection



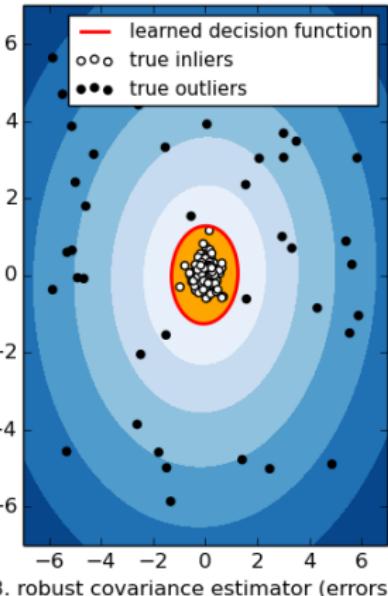
1. Isolation Forest (errors: 0)

Outlier detection



2. One-Class SVM (errors: 4)

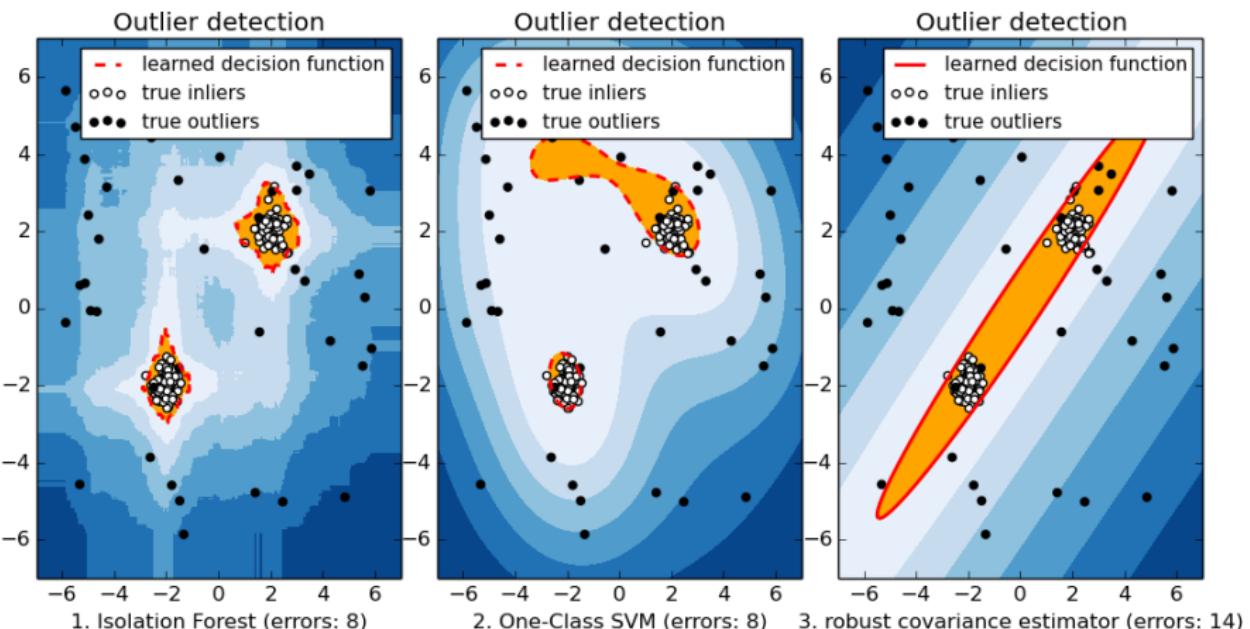
Outlier detection



3. robust covariance estimator (errors: 0)

```
n_samples_normal = 150
```

```
n_samples_outliers = 50
```

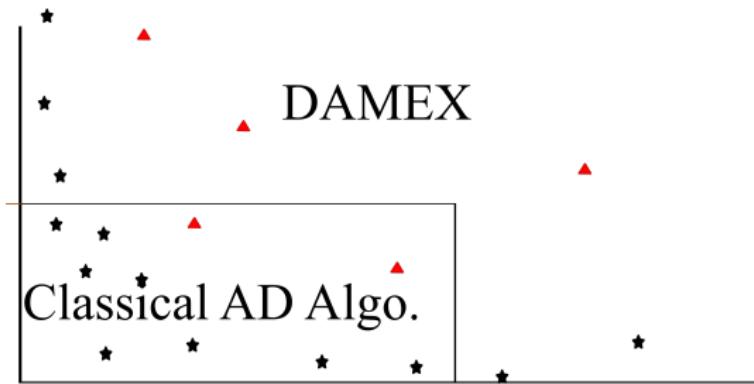


```
n_samples_normal = 150  
n_samples_outliers = 50
```

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## General idea of our work

- Extreme observations play a special role when dealing with outlying data.
- But no algorithm has **specific treatment for such multivariate extreme observations.**
- Our goal: Provide a method which can improve performance of standard AD algorithms by combining them with a **multivariate extreme analysis of the dependence structure.**



# Goal:

$$\mathbf{X} = (X_1, \dots, X_d)$$

Find the groups of features which can be large together

ex:  $\{X_1, X_2\}$ ,  $\{X_3, X_6, X_7\}$ ,  $\{X_2, X_4, X_{10}, X_{11}\}$

⇒ Characterize the extreme dependence structure

Anomalies = points which violate this structure

# Framework

- **Context**

- ▶ Random vector  $\mathbf{X} = (X_1, \dots, X_d)$
- ▶ Margins:  $X_j \sim F_j$       ( $F_j$  continuous)

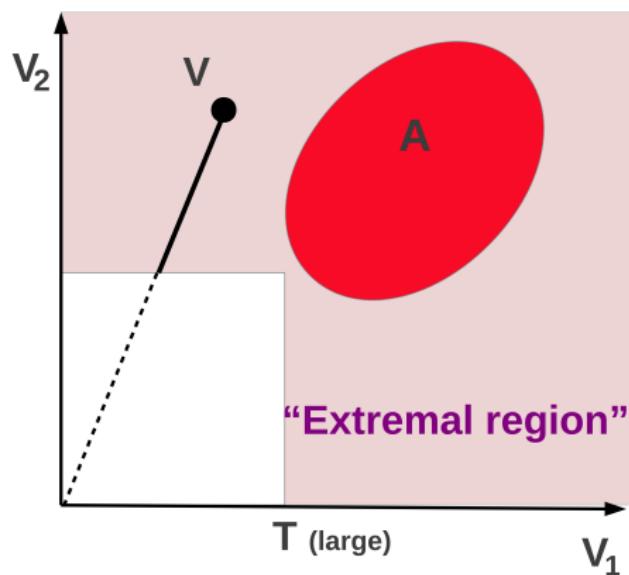
- **Preliminary step: Standardization of each marginal**

- ▶ Standard Pareto:  $V_j = \frac{1}{1-F_j(X_j)}$        $\left( \mathbb{P}(V_j \geq x) = \frac{1}{x}, \quad x \geq 1 \right)$

# Problematic

Joint extremes:  $\mathbf{V}$ 's distribution above large thresholds?

$\mathbb{P}(\mathbf{V} \in A)$ ? ( $A$  'far from the origin').



# Fundamental hypothesis and consequences

- Standard assumption: let  $A$  extreme region,

$$\mathbb{P}[\mathbf{V} \in tA] \simeq t^{-1} \mathbb{P}[\mathbf{V} \in A] \quad (\text{radial homogeneity})$$

- Formally,

**regular variation** (after standardization):

$$0 \notin \overline{A}$$

$$t\mathbb{P}[\mathbf{V} \in tA] \xrightarrow[t \rightarrow \infty]{} \mu(A), \quad \mu : \text{exponent measure}$$

$$\text{Necessarily: } \mu(tA) = t^{-1} \mu(A)$$

- $\Rightarrow$  **angular measure** on sphere  $S_{d-1}$ :  $\Phi(B) = \mu[tB, t \geq 1]$

# General model in multivariate EVT

## Model for excesses

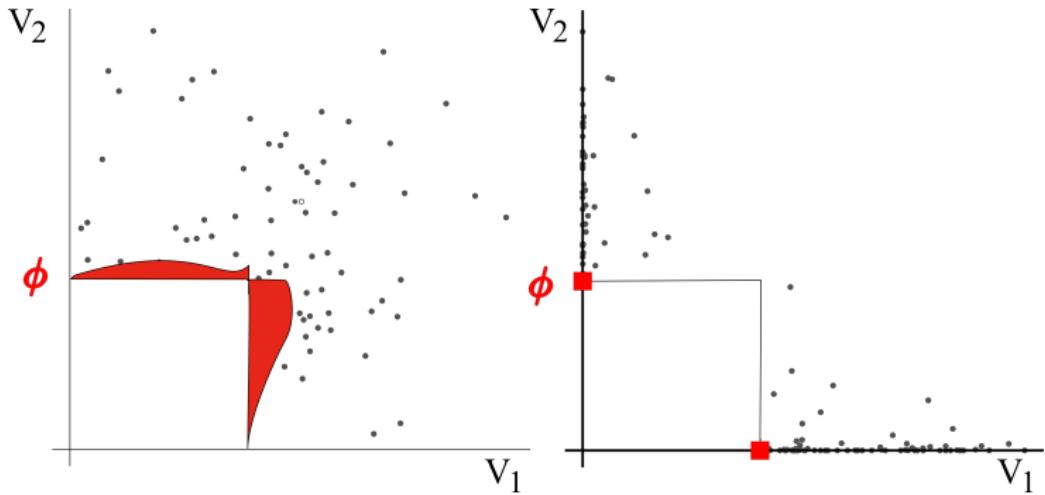
Intuitively:  $\mathbb{P}[\mathbf{V} \in A] \simeq \mu(A)$  For a large  $r > 0$  and a region  $B$  on the unit sphere:

$$\mathbb{P}\left[\|\mathbf{V}\| > r, \frac{\mathbf{V}}{\|\mathbf{V}\|} \in B\right] \sim \frac{1}{r} \Phi(B) = \mu(\{tB, t \geq r\}), r \rightarrow \infty$$

$\Rightarrow \Phi$  (or  $\mu$ ) **rules the joint distribution of extremes** (if margins are known).

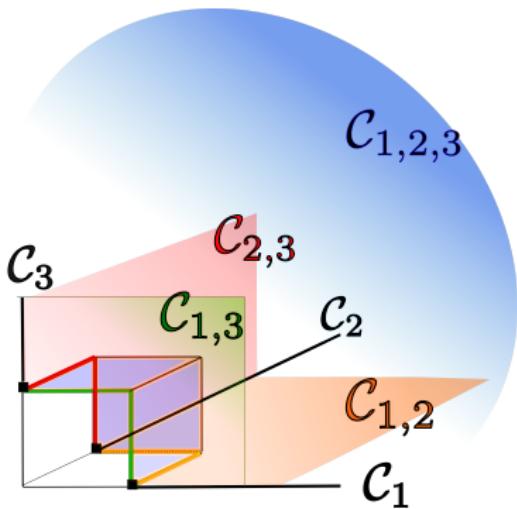
# Angular distribution

- $\Phi$  rules the joint distribution of extremes



- ▶ Asymptotic dependence:  $(V_1, V_2)$  may be large together.
- vs
- ▶ Asymptotic independence: only  $V_1$  or  $V_2$  may be large.

# General Case



- Sub-cones:  $\mathcal{C}_\alpha = \{\|\nu\| \geq 1, \nu_i > 0 \ (i \in \alpha), \nu_j = 0 \ (j \notin \alpha)\}$
- Corresponding sub-spheres:  $\{\Omega_\alpha, \alpha \subset \{1, \dots, d\}\}$   
 $(\Omega_\alpha = \mathcal{C}_\alpha \cap \mathbf{S}_{d-1})$

# Representation of extreme data

- Natural decomposition of the angular measure :

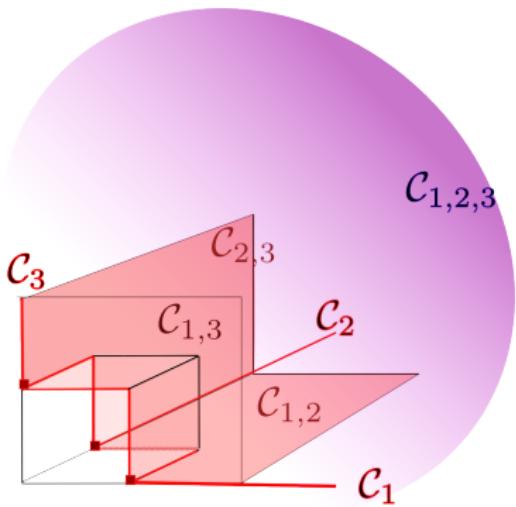
$$\Phi = \sum_{\alpha \subset \{1, \dots, d\}} \Phi_\alpha \quad \text{with } \Phi_\alpha = \Phi|_{\Omega_\alpha} \leftrightarrow \mu|_{\mathcal{C}_\alpha}$$

- $\Rightarrow$  yields a representation

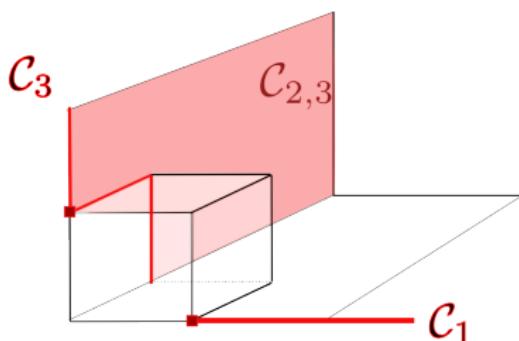
$$\begin{aligned} \mathcal{M} &= \left\{ \Phi(\Omega_\alpha) : \emptyset \neq \alpha \subset \{1, \dots, d\} \right\} \\ &= \left\{ \mu(\mathcal{C}_\alpha) : \emptyset \neq \alpha \subset \{1, \dots, d\} \right\} \end{aligned}$$

- Assumption:  $\frac{d\mu|_{\mathcal{C}_\alpha}}{d\nu_\alpha} = O(1)$ .
- Remark: Representation  $\mathcal{M}$  is linear (after non-linear transform of the data  $\mathbf{X} \rightarrow \mathbf{V}$ ).

# Sparse Representation ?



Full pattern :  
anything may happen



Sparse pattern  
( $V_1$  not large if  $V_2$  or  $V_3$  large)

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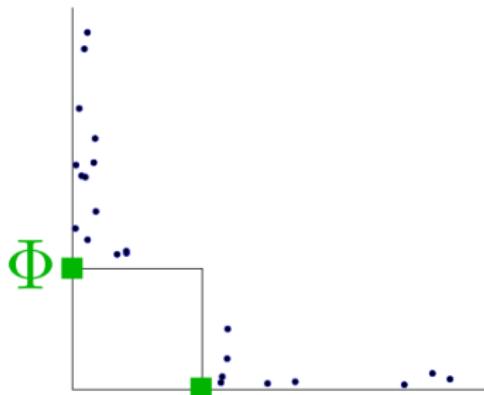
# Problem: $\mathcal{M}$ is an **asymptotic** representation

$$\mathcal{M} = \{ \Phi(\Omega_\alpha), \alpha \} = \{ \mu(\mathcal{C}_\alpha), \alpha \}$$

is the restriction of an asymptotic measure

$$\mu(A) = \lim_{t \rightarrow \infty} t \mathbb{P}[\mathbf{V} \in t A]$$

to a representative class of set  $\{\mathcal{C}_\alpha, \alpha\}$ , but only the central sub-cone has positive Lebesgue measure!

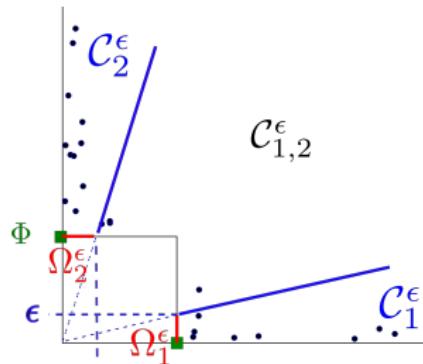


⇒ Cannot just do, for large  $t$ :

$$\Phi(\Omega_\alpha) = \mu(\mathcal{C}_\alpha) \simeq t \widehat{\mathbb{P}}(t \mathcal{C}_\alpha)$$

# Solution

**Fix  $\epsilon > 0$ . Affect data  $\epsilon$ -close to an edge, to that edge.**



$$\Omega_\alpha \rightarrow \Omega_\alpha^\epsilon = \{v \in \mathbf{S}_{d-1} : v_i > \epsilon \ (j \in \alpha), v_j \leq \epsilon \ (j \notin \alpha)\}.$$

$$\mathcal{C}_\alpha \rightarrow \mathcal{C}_\alpha^\epsilon = \{t \Omega_\alpha^\epsilon, t \geq 1\}$$

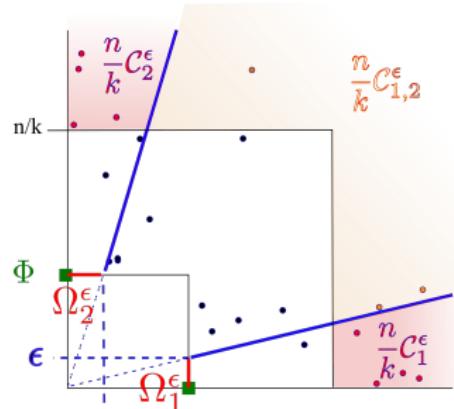
New partition of  $\mathbf{S}_{d-1}$ , compatible with non asymptotic data.

$$\hat{V}_i^j = \frac{1}{1 - \hat{F}_j(X_i^j)} \text{ with } \hat{F}_j(X_i^j) = \frac{\text{rank}(X_i^j) - 1}{n}$$

$\Rightarrow$  get an natural estimate of  
 $\Phi(\Omega_\alpha)$

$$\hat{\Phi}(\Omega_\alpha) := \frac{n}{k} \mathbb{P}_n(\hat{V} \in \frac{n}{k} \mathcal{C}_\alpha^\epsilon)$$

( $\frac{n}{k}$  large,  $\epsilon$  small)



$\Rightarrow$  we obtain

$$\widehat{\mathcal{M}} := \{ \hat{\Phi}(\Omega_\alpha), \alpha \}$$

## Theorem

*There is an absolute constant  $C > 0$  such that for any  $n > 0$ ,  $k > 0$ ,  $0 < \epsilon < 1$ ,  $\delta > 0$  such that  $0 < \delta < e^{-k}$ , with probability at least  $1 - \delta$ ,*

$$\|\widehat{\mathcal{M}} - \mathcal{M}\|_\infty \leq Cd \left( \sqrt{\frac{1}{\epsilon k} \log \frac{d}{\delta}} + Md\epsilon \right) + \text{bias}(\epsilon, k, n),$$

### Comments:

- $C$ : depends on  $M = \sup(\text{density on subfaces})$
- Existing litterature (for spectral measure) **Einmahl Segers 09, Einmahl et.al. 01**

$$d = 2.$$

asymptotic behaviour, rates in  $1/\sqrt{k}$ .

**Here:**  $1/\sqrt{k} \rightarrow 1/\sqrt{\epsilon k} + \epsilon$ . Price to pay for biasing our estimator with  $\epsilon$ .

## Theorem's proof

### ① Maximal deviation on VC-class:

$$\sup_{x \geq \epsilon} |\mu_n - \mu|([x, \infty]) \leq Cd \sqrt{\frac{2}{k} \log \frac{d}{\delta}} + \text{bias}(\epsilon, k, n)$$

**Tools:** Vapnik-Chervonenkis inequality adapted to small probability sets: bounds in  $\sqrt{p} \sqrt{\frac{1}{n} \log \frac{1}{\delta}}$

On the VC class  $\{[\frac{n}{k}x, \infty], x \geq \epsilon\}$

## Theorem's proof

- ① **Maximal deviation on VC-class:**
- ② **Decompose error:**

$$|\mu_n(\mathcal{C}_\alpha^\epsilon) - \mu(\mathcal{C}_\alpha)| \leq \underbrace{|\mu_n - \mu|(\mathcal{C}_\alpha^\epsilon)}_A + \underbrace{|\mu(\mathcal{C}_\alpha^\epsilon) - \mu(\mathcal{C}_\alpha)|}_B$$

- ▶ A : First step.
- ▶ B : density on  $\mathcal{C}_\alpha^\epsilon \times Lebesgue$  : small

## Algorithm

DAMEX in  $O(dn \log n)$

**Input:** parameters  $\epsilon > 0$ ,  $k = k(n)$ ,

- ① Standardize via marginal rank-transformation:

$$\hat{V}_i := (1/(1 - \hat{F}_j(X_i^j)))_{j=1,\dots,d}.$$

- ② Assign to each  $\hat{V}_i$  the cone  $\frac{n}{k}\mathcal{C}_\alpha^\epsilon$  it belongs to.

- ③  $\Phi_n^{\alpha, \epsilon} := \hat{\Phi}(\Omega_\alpha) = \frac{n}{k}\mathbb{P}_n(\hat{V} \in \frac{n}{k}\mathcal{C}_\alpha^\epsilon)$  the estimate of the  $\alpha$ -mass of  $\Phi$ .

**Output:** (sparse) representation of the dependence structure

$$\widehat{\mathcal{M}} := (\Phi_n^{\alpha, \epsilon})_{\alpha \subset \{1, \dots, d\}, \Phi_n^{\alpha, \epsilon} > \Phi_{\min}}$$

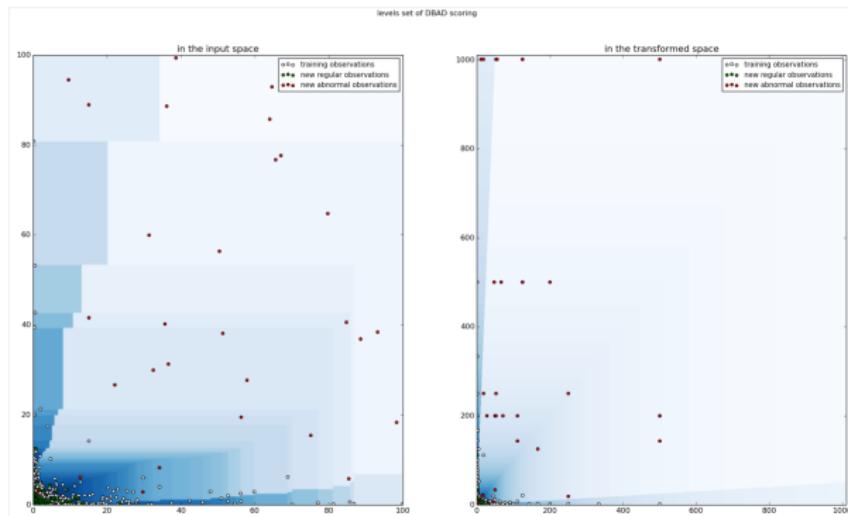
# Application to Anomaly Detection

After standardization of marginals:  $\mathbb{P}[R > r, \mathbf{W} \in \mathcal{B}] \simeq \frac{1}{r} \Phi(B)$

→ scoring function =  $\Phi_n^\epsilon \times 1/r$ :

$$s_n(\mathbf{x}) := (1/\|\hat{T}(\mathbf{x})\|_\infty) \sum_{\alpha} \Phi_n^{\alpha, \epsilon} \mathbb{1}_{\hat{T}(\mathbf{x}) \in \mathcal{C}_\alpha^\epsilon}.$$

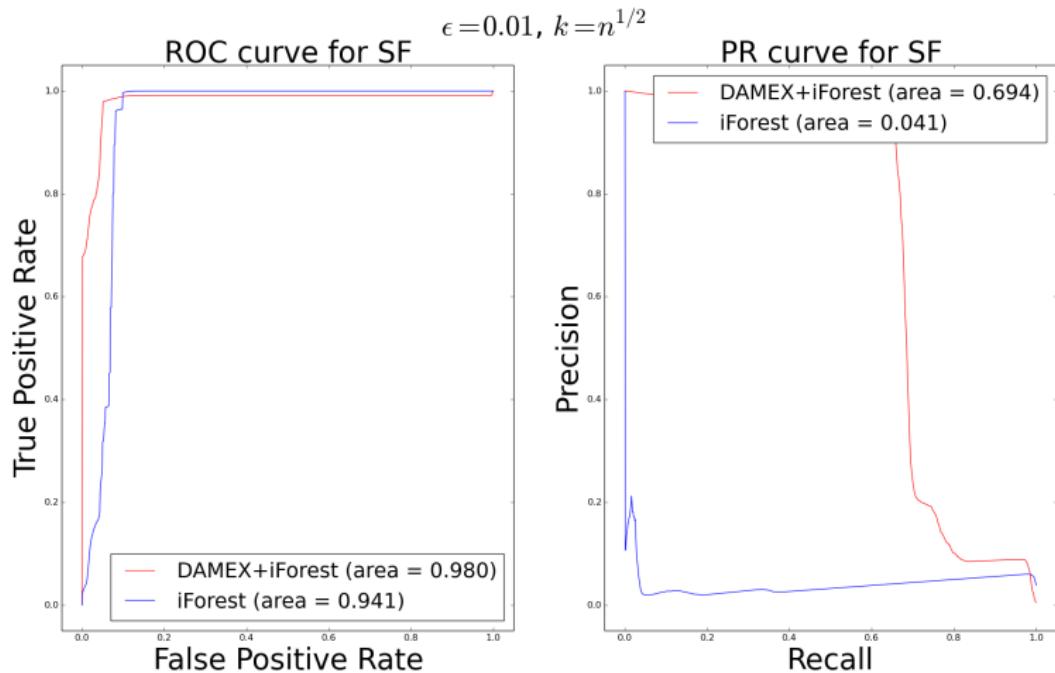
where  $T : \mathbf{X} \mapsto \mathbf{V}$       ( $V_j = \frac{1}{1-F_j(X_j)}$ )



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	number of samples	number of features
shuttle	85849	9
forestcover	286048	54
SA	976158	41
SF	699691	4
http	619052	3
smtp	95373	3

Table: Datasets characteristics



**Figure:** ROC and PR curve on SF dataset

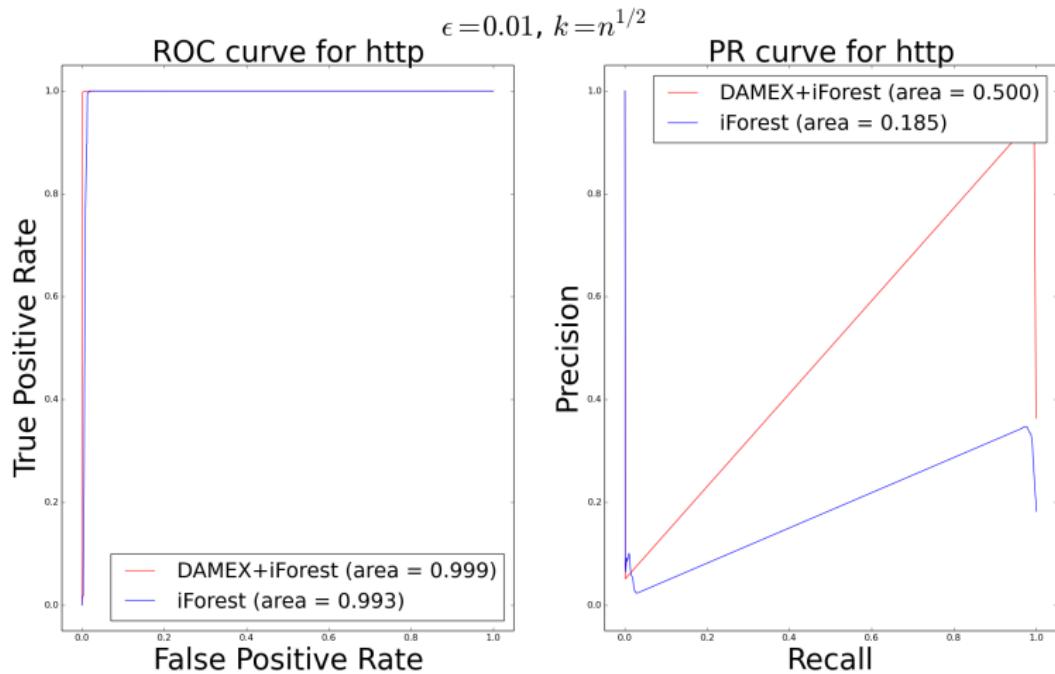


Figure: ROC and PR curve on http dataset

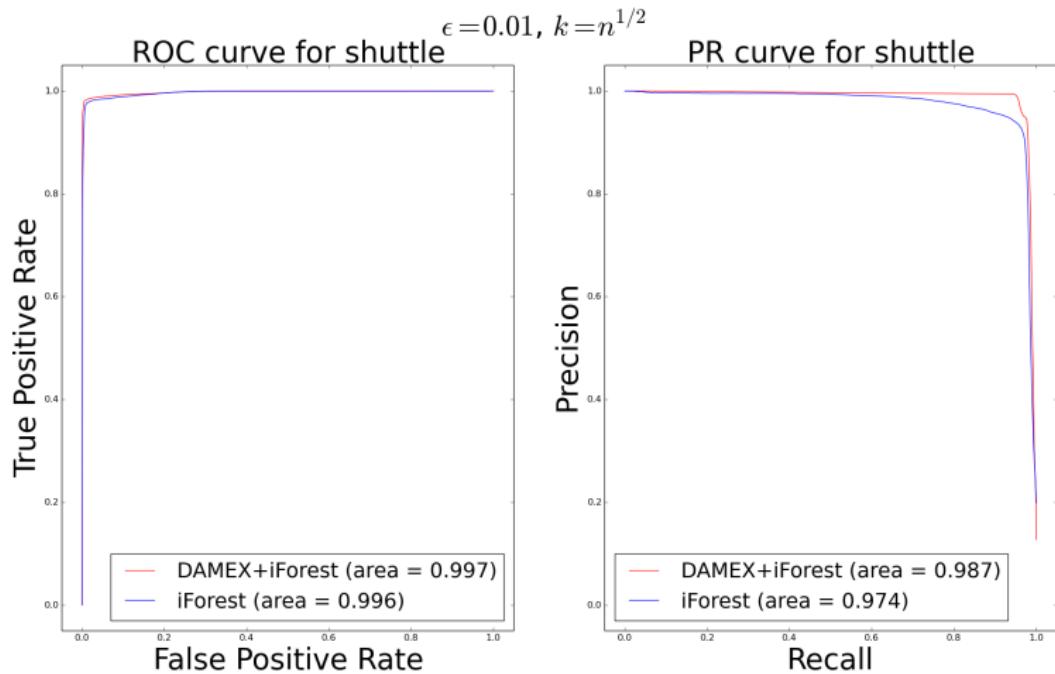
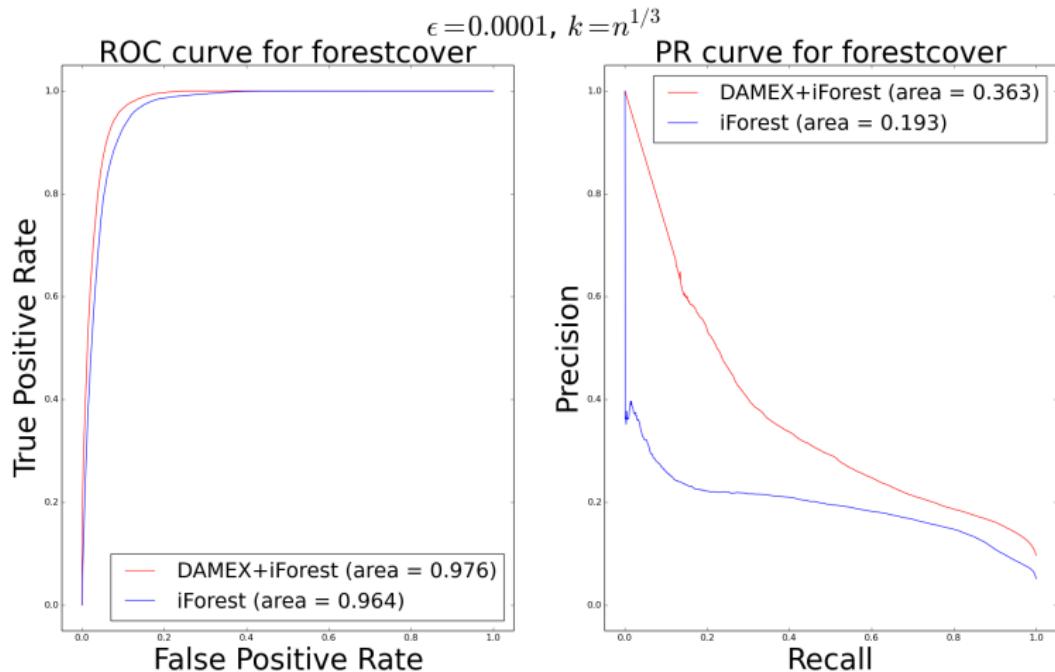
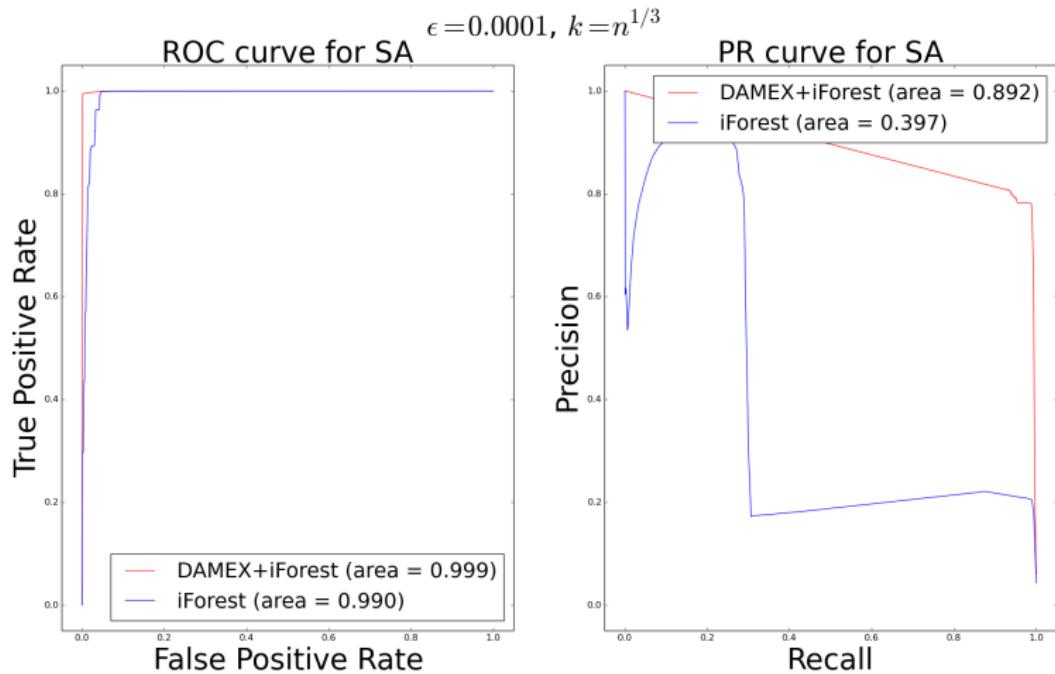


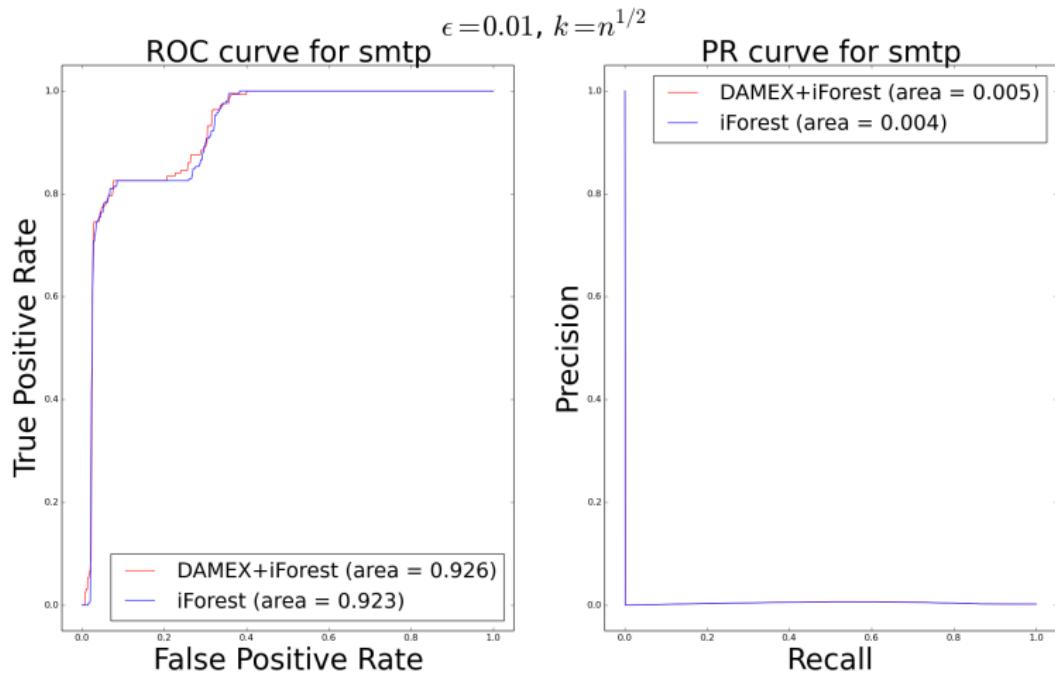
Figure: ROC and PR curve on shuttle dataset



**Figure:** ROC and PR curve on forestcover dataset



**Figure:** ROC and PR curve on SA dataset



**Figure:** ROC and PR curve on smtp dataset

Thank you !

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